

### Introduction

- The last couple of lectures looked at techniques to handle *uncertainty* 
  - Bayesian networks
- The formalism is *static*, and so has limited ability to handle changing information.
- Lots of reasoning tasks involve a dynamic world
  - Monitoring a patient
  - Tracking an airplane
  - Identifying the location of a robot
- This week we'll look at models that can handle such dynamic situations.
  - Based on Bayesian networks

## States and observations

- The approach we'll look at considers the world to be a series of *time slices*.
- Each slice contains some variables:
  - The set  $X_t$  which we can't observe; and
  - The set  $\mathbf{E}_t$  which we can observe.
- At a given point in time we have an observation  $\mathbf{E}_t = \mathbf{e}_t$ .
- What would be an example?

- Consider you live and work in some location without a window
  - Not so hard to imagine when you know the GC



- You want to know whether it is raining.
- Your only information is looking at whether somebody who comes into your location each morning is carrying an umbrella.

- Each day is one value of *t*.
- $\mathbf{E}_t$  contains the single variable  $U_t$  (or  $Umbrella_t$ ).
  - Is the person carrying an umbrella?
- $X_t$  contains the single variable  $R_t$  (or  $Rain_t$ )
  - Is it raining?

- State sequence starts at t = 0, and the interval between slices in general depends on the problem.
  - Here it is one day
  - In robot localization it is pretty arbitrary
- First piece of evidence arrives at t = 1
- So, the umbrella world is:

$$R_0, R_1, R_2, \dots$$

$$U_1, U_2, U_3, \dots$$

• a:b means the sequence of integers from a to b, so that  $U_{2:4}$  is the sequence:

$$U_2, U_3, U_4$$

## Transition and sensor models

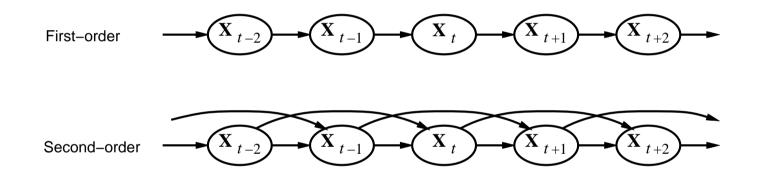
- We need to add two components to this backbone:
  - How the world evolves
     Transition model
  - What the evidence tells us Sensor model
- The transition model tells us:

$$\mathbf{P}(\mathbf{X}_t|\mathbf{X}_{0:t-1})$$

what the probability is that it is raining today given the weather every previous day for as long as records have existed.

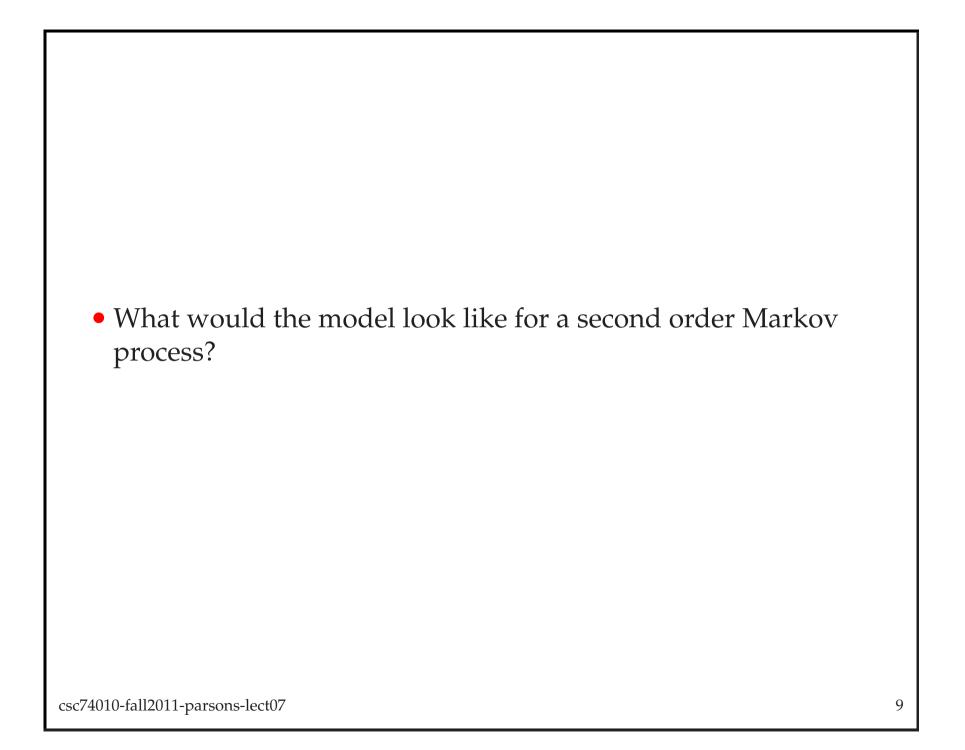
• What is the problem with this model?

- Luckily Professor Markov helps us out again.
- Make a *Markov assumption* that the value of the current state depends only on a finite fixed number of previous states.



• We commonly assume a *first order* Markov process, where the current state depends only on the previous state:

$$\mathbf{P}(\mathbf{X}_t|\mathbf{X}_{0:t-1}) = \mathbf{P}(\mathbf{X}_t|\mathbf{X}_{t-1})$$



• What would the model look like for a second order Markov process?

$$\mathbf{P}(\mathbf{X}_t|\mathbf{X}_{0:t-1}) = \mathbf{P}(\mathbf{X}_t|\mathbf{X}_{t-2},\mathbf{X}_{t-1})$$

• Even with the Markov assumption we have a potentially infinite set of conditional probabilities.

$$P(X_1|X_0), P(X_2|X_1), P(X_3|X_2)...$$

- Usually circumvent this by assuming a *stationary process* 
  - The model doesn't change
  - But the state itself can
- Thus we only have one, general  $P(X_t|X_{t-1})$

### An aside

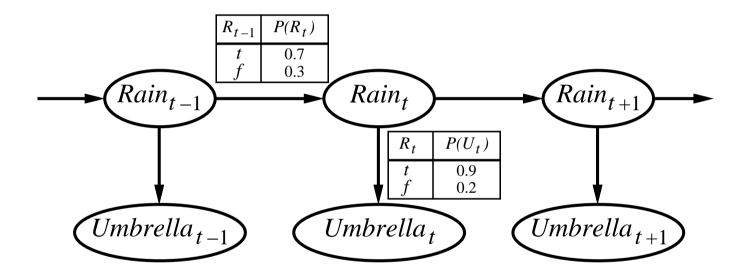
- First-order Markov assumption not exactly true in real world!
- Possible fixes:
  - 1. *Increase order* of Markov process
  - 2. Augment state, e.g., add Tempt, Pressuret
- Example: robot motion.
  - Augment position and velocity with  $Battery_t$
- Example: umbrella world.
  - Augment state with  $Season_t$  and/or  $Pressure_t$

Sensor model

- The evidence variables  $\mathbf{E}_t$  could depend on lots of previous variables.
- But we will assume the state is constructed in such a way that evidence only depends on the current state.
- A Markov assumption for the sensor model:

$$\mathbf{P}(\mathbf{E}_t|\mathbf{X}_{0:t-1},\mathbf{E}_{0:t-1}) = \mathbf{P}(\mathbf{E}_t|\mathbf{X}_t)$$

• Here are the sensor and observation models for the umbrella world:



• As for previous Bayesian networks, arrows run from causes to effects.

• Also need to say how things get started:

 $\mathbf{P}(\mathbf{X}_0)$ 

The prior probability over the state

• With this, we can then compute the complete joint probability over all the time slices:

$$\mathbf{P}(\mathbf{X}_{0:t}, \mathbf{E}_{1:t}) = \mathbf{P}(\mathbf{X}_0) \prod_{i=1}^{t} \mathbf{P}(\mathbf{X}_i | \mathbf{X}_{i-1}) \mathbf{P}(\mathbf{E}_i | \mathbf{X}_i)$$

• As we know from before, this is sufficient to compute anything we want.

#### Inference tasks

- What kinds of thing can we do with the model?
- *Filtering*:  $P(X_t|e_{1:t})$ 
  - determine belief state—input to the decision process of a rational agent
- *Prediction*:  $P(X_{t+k}|e_{1:t})$  for k>0
  - Evaluation of possible action sequences, like filtering without the evidence
- *Smoothing*:  $P(\mathbf{X}_k | \mathbf{e}_{1:t})$  for  $0 \le k < t$ 
  - Better estimate of past states, essential for learning
- Most likely explanation:  $\arg \max_{\mathbf{X}_{1:t}} P(\mathbf{x}_{1:t}|\mathbf{e}_{1:t})$ 
  - Speech recognition, decoding with a noisy channel

Filtering

• A good algorithm for filtering will maintain a current state estimate and update it at each point.

$$\mathbf{P}(\mathbf{X}_{t+1}|\mathbf{e}_{1:t+1}) = f(\mathbf{P}(\mathbf{X}_t|\mathbf{e}_{1:t}),\mathbf{e}_{t+1})$$

- Saves recomputation.
- It turns out that this is easy enough to come up with.

• We rearrange the formula for:

$$\mathbf{P}(\mathbf{X}_{t+1}|\mathbf{e}_{1:t+1})$$

• First, we divide up the evidence:

$$\mathbf{P}(\mathbf{X}_{t+1}|\mathbf{e}_{1:t+1}) = \mathbf{P}(\mathbf{X}_{t+1}|\mathbf{e}_{1:t},\mathbf{e}_{t+1})$$

• Then we apply Bayes rule, remembering the use of the normalization factor  $\alpha$ .

$$\mathbf{P}(\mathbf{X}_{t+1}|\mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{t+1}|\mathbf{X}_{t+1},\mathbf{e}_{1:t})\mathbf{P}(\mathbf{X}_{t+1}|\mathbf{e}_{1:t})$$

• And after that we use the Markov assumption on the sensor model:

$$\mathbf{P}(\mathbf{X}_{t+1}|\mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{t+1}|\mathbf{X}_{t+1})\mathbf{P}(\mathbf{X}_{t+1}|\mathbf{e}_{1:t})$$

• The result of this assumption is to make that first term on the right hand side ignore all the evidence — the probability of the observation at t + 1 only depends on the value of  $X_{t+1}$ .

• Let's look at that expression some more:

$$\mathbf{P}(\mathbf{X}_{t+1}|\mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{t+1}|\mathbf{X}_{t+1})\mathbf{P}(\mathbf{X}_{t+1}|\mathbf{e}_{1:t})$$

- The first term on the right updates with the new evidence and the second term on the right is a one step prediction from the evidence up to t to the state at t + 1.
- Next we condition on the current state P(X):

$$\mathbf{P}(\mathbf{X}_{t+1}|\mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{t+1}|\mathbf{X}_{t+1}) \sum_{\mathbf{X}_t} \mathbf{P}(\mathbf{X}_{t+1}|\mathbf{x}_t,\mathbf{e}_{1:t}) P(\mathbf{x}_t|\mathbf{e}_{1:t})$$

• Finally, we apply the Markov assumption again:

$$\mathbf{P}(\mathbf{X}_{t+1}|\mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{t+1}|\mathbf{X}_{t+1}) \sum_{\mathbf{X}_t} \mathbf{P}(\mathbf{X}_{t+1}|\mathbf{X}_t) P(\mathbf{X}_t|\mathbf{e}_{1:t})$$

• We'll call the bit on the right  $\mathbf{f}_{1:t}$ 

- $\mathbf{f}_{1:t}$  gives us the required recursive update.
  - The probability distribution over the state variables at t + 1 is a function of the transition model, the sensor model, and what we know about the state at time t.
- Space and time constant, independent of *t*.
- This allows a limited agent to compute the current distribution for any length of sequence.

# Filtering the umbrella example

- The prior is (0.5, 0.5).
- We can first predict whether it will rain on day 1 given what we already know:

$$\mathbf{P}(\mathbf{R}_1) = \sum_{r_0} \mathbf{P}(R_1|r_0) P(r_0)$$

$$= \langle 0.7, 0.3 \rangle \times 0.5 + \langle 0.3, 0.7 \rangle \times 0.5$$

$$= \langle 0.5, 0.5 \rangle$$

• As we should expect, this just gives us the prior — that is the probability of rain when we don't have any evidence.

• However, we have observed the umbrella, so that  $U_1 = true$ , and we can update using the sensor model:

$$\mathbf{P}(\mathbf{R}_1|U_1) = \alpha \mathbf{P}(u_1|R_1)\mathbf{P}(R_1)$$

$$= \alpha \langle 0.9, 0.2 \rangle \langle 0.5, 0.5 \rangle$$

$$= \alpha \langle 0.45, 0.1 \rangle$$

$$\approx \langle 0.818, 0.182 \rangle$$

• So, since umbrella is strong evidence for rain, the probability of rain is much higher once we take the observation into account.

• We can then carry out the same computation for Day 2, first predicting whether it will rain on day 1 given what we already know:

$$\mathbf{P}(\mathbf{R}_{2}|u_{1}) = \sum_{r_{1}} \mathbf{P}(R_{2}|r_{1})P(r_{1}|u_{1}) 
= \langle 0.7, 0.3 \rangle \times 0.818 + \langle 0.3, 0.7 \rangle \times 0.182 
\approx \langle 0.627, 0.373 \rangle$$

• So even without evidence of rain on the second day there is a higher probability of rain than the prior because rain tends to follow rain.

(In this model rain tends to persist.)

• Then we can repeat the evidence update,  $u_2$  ( $U_2 = true$ ), so:

$$\mathbf{P}(\mathbf{R}_{2}|u_{1}, u_{2}) = \alpha \mathbf{P}(u_{2}|R_{2})\mathbf{P}(R_{2}|u_{1})$$

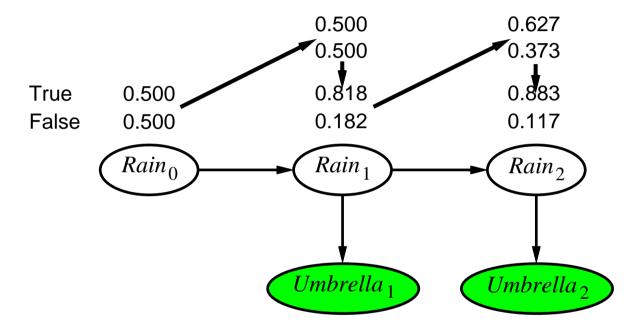
$$= \alpha \langle 0.9, 0.2 \rangle \langle 0.627, 0.373 \rangle$$

$$= \alpha \langle 0.565, 0.075 \rangle$$

$$\approx \langle 0.883, 0.117 \rangle$$

• So, the probability of rain increases again, and is higher than on day 1.

• Put more succinctly:



• We can think of the calculation as messages passed along the chain.

Prediction

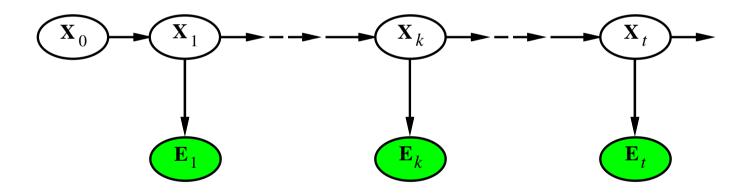
Prediction is filtering without new evidence.

$$\mathbf{P}(\mathbf{X}_{t+1}|\mathbf{e}_{1:t+1}) = \sum_{\mathbf{X}_t} \mathbf{P}(\mathbf{X}_{t+1}|\mathbf{X}_t,\mathbf{e}_{1:t}) P(\mathbf{X}_t|\mathbf{e}_{1:t})$$

- Given the current state, what does the future bring?
- Just need the first part of the calculation we did above (in each step we first predicted and then updated with evidence).

# Smoothing

• Smoothing is computing the distribution over past states given evidence up to the present.



• Want the probability over all states k,  $0 \le k < t$ .

- Break the computation into two pieces, evidence from 0 to k and evidence from k + 1 to t.
- Proceeding just as before:

$$\mathbf{P}(\mathbf{X}_{k}|\mathbf{e}_{1:t}) = \mathbf{P}(\mathbf{X}_{k}|\mathbf{e}_{1:k},\mathbf{e}_{k+1:t})$$

$$= \alpha \mathbf{P}(\mathbf{X}_{k}|\mathbf{e}_{1:k})\mathbf{P}(\mathbf{e}_{k+1:t}|\mathbf{X}_{k},\mathbf{e}_{1:k})$$

$$= \alpha \mathbf{P}(\mathbf{X}_{k}|\mathbf{e}_{1:k})\mathbf{P}(\mathbf{e}_{k+1:t}|\mathbf{X}_{k})$$

$$= \alpha \mathbf{f}_{1:k}\mathbf{b}_{k+1:t}$$

- **f** is a "forward" message, computed just as we did for the filtering case.
- **b** is a backward message.

• To compute the backwards message we condition on  $X_{k+1}$ :

$$\mathbf{P}(\mathbf{e}_{k+1:t}|\mathbf{X}_k) = \sum_{\mathbf{X}_{k+1}} \mathbf{P}(\mathbf{e}_{k+1:t}|\mathbf{X}_k, \mathbf{X}_{k+1}) \mathbf{P}(\mathbf{X}_{k+1}|\mathbf{X}_k)$$

Then apply conditional independence:

$$\mathbf{P}(\mathbf{e}_{k+1:t}|\mathbf{X}_k) = \sum_{\mathbf{X}_{k+1}} P(\mathbf{e}_{k+1:t}|\mathbf{X}_{k+1}) \mathbf{P}(\mathbf{X}_{k+1}|\mathbf{X}_k)$$

Condition again:

$$\mathbf{P}(\mathbf{e}_{k+1:t}|\mathbf{X}_k) = \sum_{\mathbf{X}_{k+1}} P(\mathbf{e}_{k+1}|\mathbf{X}_{k+1}) P(\mathbf{e}_{k+2:t}|\mathbf{X}_{k+1}) \mathbf{P}(\mathbf{X}_{k+1}|\mathbf{X}_k)$$

• The first and third terms here come from the model, the second term is the bit we compute recursively.

• So in:

$$\mathbf{f}_{1:k}\mathbf{b}_{k+1:t}$$

there are two recursive components.

- We have a forward component from 1 to k, and a backward component from t to k.
  - The backward component is initialized with:

$$\mathbf{P}(\mathbf{e}_{t+1:t}|\mathbf{X}_t) = \mathbf{P}(\ |\mathbf{X}_t) = 1$$

(the probability of observing the set of no observations is always 1).

• Consider the umbrella world on day 1. This time we update with information about umbrellas on day 1 and day 2:

$$\mathbf{P}(R_1|u_1,u_2) = \alpha \mathbf{P}(R_1|u_1)\mathbf{P}(u_2|R_1)$$

We know that the first of these terms is (0.818, 0.182) from before.

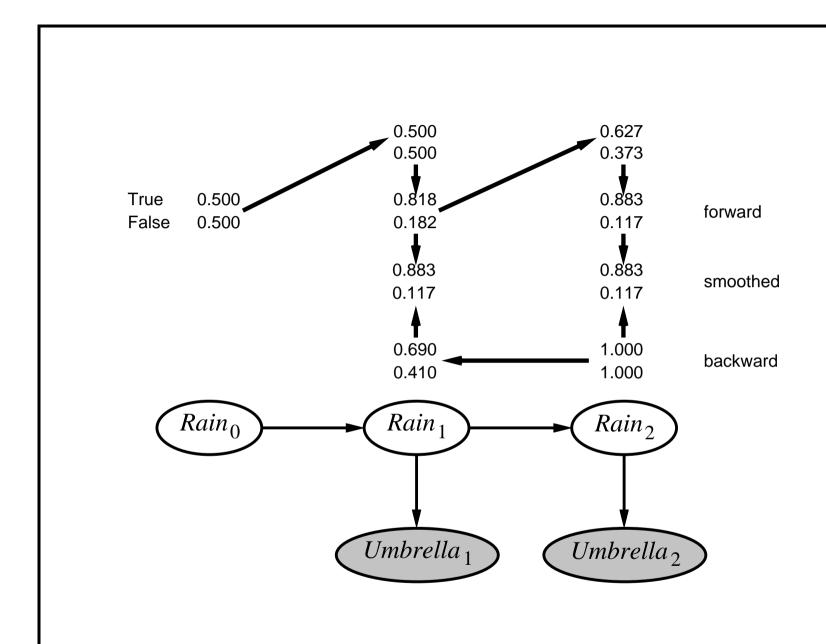
• The second term we can compute:

$$\mathbf{P}(u_2|R_1) = \sum_{r_2} P(u_2|r_2) P(|r_2) \mathbf{P}(r_2|R_1) 
= (0.9 \times 1 \times \langle 0.70.3 \rangle) + (0.2 \times 1 \times \langle 0.3, 0.7 \rangle) 
= \langle 0.69, 0.41 \rangle$$

• Plugging this back into the expression we started with gives:

$$\mathbf{P}(R_1|u_1, u_2) = \alpha \langle 0.818, 0.182 \rangle \times \langle 0.69, 0.41 \rangle$$
  
=  $\alpha \langle 0.56, 0.075 \rangle$   
 $\approx \langle 0.882, 0.118 \rangle$ 

• Again we can think of this as message passing (next slide)



• The smoothed probability of rain on day 1 is *higher* than the filtered estimate because the umbrella on day 2 makes it more likely to have rained on day 1.



• Again these updates use constant time and space.

# Most likely sequence

- Most likely sequence ≠ sequence of most likely states!!!!
- Most likely path to each  $\mathbf{x}_{t+1}$  is most likely path to *some*  $\mathbf{x}_t$  plus one more step

$$\max_{\mathbf{X}_1...\mathbf{X}_t} \mathbf{P}(\mathbf{x}_1, \dots, \mathbf{x}_t, \mathbf{X}_{t+1} | \mathbf{e}_{1:t+1})$$

$$= \mathbf{P}(\mathbf{e}_{t+1} | \mathbf{X}_{t+1}) \max_{\mathbf{X}_t} \left( \mathbf{P}(\mathbf{X}_{t+1} | \mathbf{x}_t) \max_{\mathbf{X}_1...\mathbf{X}_{t-1}} P(\mathbf{x}_1, \dots, \mathbf{x}_{t-1}, \mathbf{x}_t | \mathbf{e}_{1:t}) \right)$$

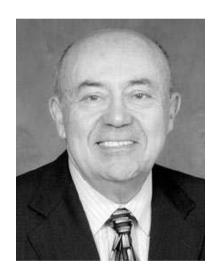
• Identical to filtering, except  $\mathbf{f}_{1:t}$  replaced by:

$$\mathbf{m}_{1:t} = \max_{\mathbf{X}_1...\mathbf{X}_{t-1}} \mathbf{P}(\mathbf{X}_1,\ldots,\mathbf{X}_{t-1},\mathbf{X}_t|\mathbf{e}_{1:t}),$$

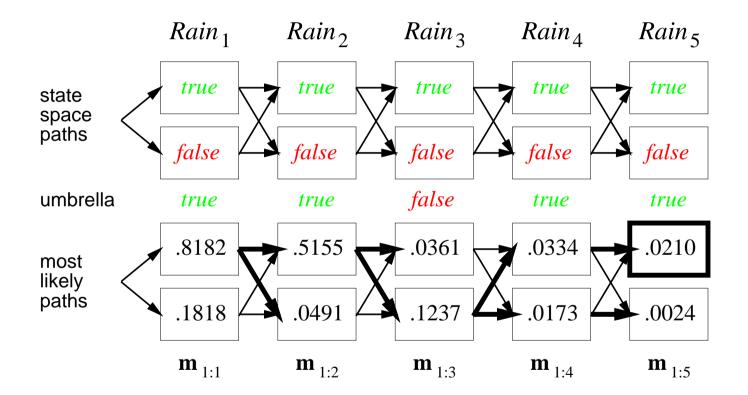
•  $\mathbf{m}_{1:t}(i)$  gives the probability of the most likely path to state *i*.

• Update has sum replaced by max, giving the *Viterbi algorithm*:

$$\mathbf{m}_{1:t+1} = \mathbf{P}(\mathbf{e}_{t+1}|\mathbf{X}_{t+1}) \max_{\mathbf{X}_t} (\mathbf{P}(\mathbf{X}_{t+1}|\mathbf{x}_t)\mathbf{m}_{1:t})$$



Andrew Viterbi



- Bold state is the most likely state at day 5.
- For each state you can tell its best predecessor (bold arrow).
- So having computed the state values, can easily read off the most likely sequence.

Where we are?

- We have a general approach to all the inference problems without thinking much about the specific details of the models.
- When we get specific, we find we can solve several common classes of problem.

#### Hidden Markov models

- HMMs have  $X_t$  as a single, discrete variable
  - Usually  $\mathbf{E}_t$  is too)
- Domain of  $X_t$  is  $\{1,\ldots,S\}$
- Transition matrix  $\mathbf{T}_{ij} = P(X_t = j | X_{t-1} = i)$

$$-\text{ e.g.,} \begin{pmatrix} 0.7 & 0.3 \\ 0.3 & 0.7 \end{pmatrix}$$

• *Sensor matrix*  $\mathbf{O}_t$  for each time step, diagonal elements  $P(e_t|X_t=i)$ 

- e.g., with 
$$U_1 = true$$
,  $\mathbf{O}_1 = \begin{pmatrix} 0.9 & 0 \\ 0 & 0.2 \end{pmatrix}$ 

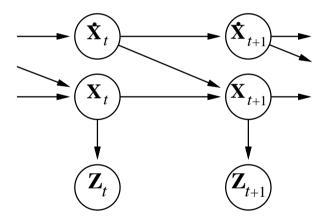
• Forward and backward messages as column vectors:

$$\mathbf{f}_{1:t+1} = \alpha \mathbf{O}_{t+1} \mathbf{T}^{\top} \mathbf{f}_{1:t}$$
  
 $\mathbf{b}_{k+1:t} = \mathbf{TO}_{k+1} \mathbf{b}_{k+2:t}$ 

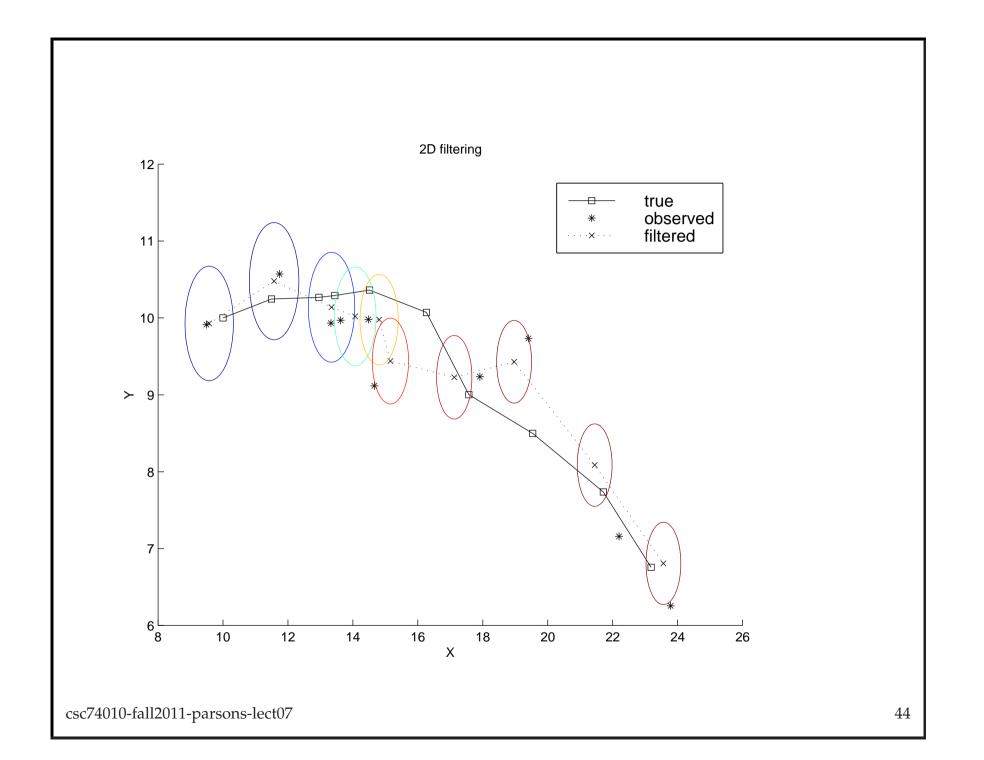
- Forward-backward algorithm needs time  $O(S^2t)$  and space O(St)
- Matrix description points the way to easy implementation.

## Kalman filters

- Modelling systems described by a set of continuous variables
  - Robot tracking — $X_t = X, Y, \dot{X}, \dot{Y}$
  - Airplanes, ecosystems, economies, chemical plants, etc



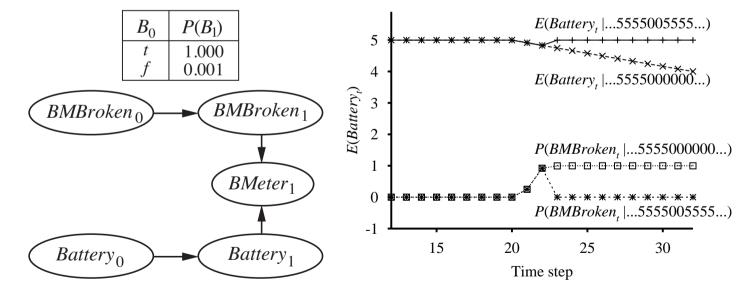
- *Z* is observation.
- Gaussian prior, linear Gaussian transition model and sensor model



- Sparse dependencies ⇒ exponentially fewer parameters;
  - e.g., 20 state variables, three parents each
- DBN has  $20 \times 2^3 = 160$  parameters, HMM has  $2^{20} \times 2^{20} \approx 10^{12}$

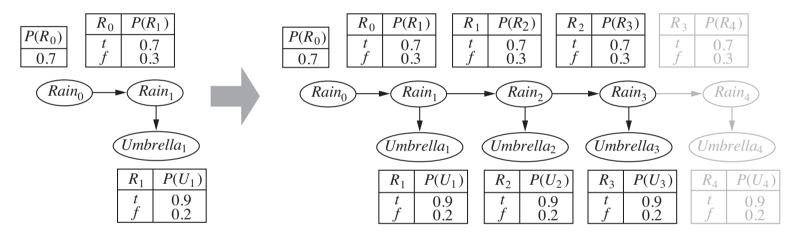
### DBNs vs Kalman filters

- Every Kalman filter model is a DBN, but few DBNs are KFs;
  - Real world requires non-Gaussian posteriors
- What's the battery charge?



### Inference in DBNs

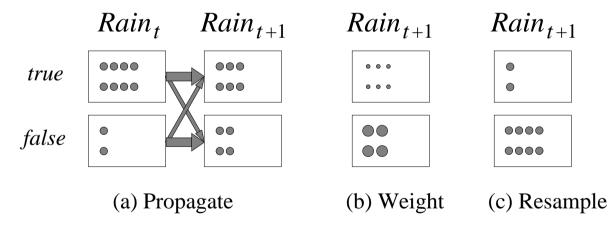
• There are exact algorithms for DBNs:



- "Unroll" DBN to create a static Bayesian network and use standard approaches.
- But time and space complexity is exponential.
- However, there are some good approximation algorithms.

#### Particle filters

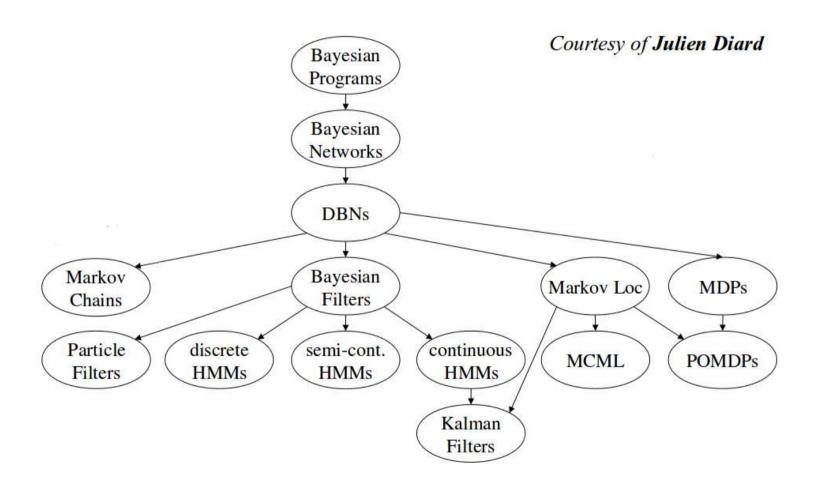
- Technique for approximate solution of a DBN.
- Basic idea: ensure that the population of samples ("particles") tracks the high-likelihood regions of the state-space
- Replicate particles proportional to likelihood for  $\mathbf{e}_t$



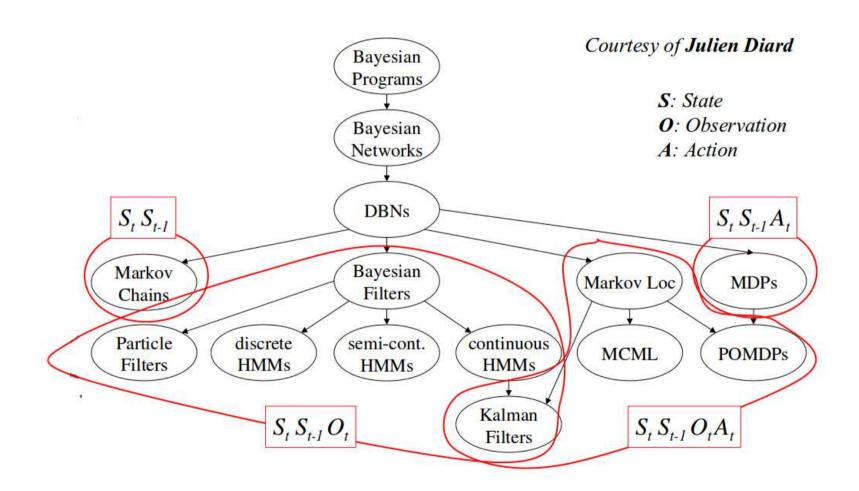
• Time/space complexity linear in the number of particles.

- Widely used for tracking nonlinear systems, esp. in vision
- Also used for simultaneous localization and mapping in mobile robots
  - 10<sup>5</sup>-dimensional state space
- Approximation error of particle filtering remains bounded over time.
- At least empirically—theoretical analysis is difficult.

A general classification of dynamic probabilistic models



A general classification of dynamic probabilistic models



# Summary

- This class moved from the static view of the world encorporated in Bayesian networks to something more dynamic.
- Dynamic Bayesian networks.
- We looked at the general types of inference possible in DBNs and showed how the necessary computations could be done.
- We also looked at some specific classes of problem that can be captured by DBNs.